

# PARALLEL PROCESSING OF 0—1 SEQUENCES

By

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## §. 1. Introduction

This paper is devoted to the investigation of parallel processing of binary sequences.

Processing proceeds in the points of time 1, 2, ..., and it is sequential for every sequence.

The sequences have priorities: processing algorithm considers the elements when choosing them for processing at every point of time in the order first, second, ..., last sequence.

The processing in every point of time is characterized by using the vector of the first non-processed elements of the sequences.

In §. 2. the formal description of the considered problem is given, later (§. 3.) the limit distribution of the vectors of the first non-processed elements is computed.

Finally (§. 4.) the limit distribution is used to get the processing speed.

## §. 2. Formulation of the problem

Let

$$(2.1) \quad \begin{array}{l} F_1 = f_{11}, f_{12}, \dots \\ \vdots \\ F_r = f_{r1}, f_{r2}, \dots \end{array}$$

$r$  ( $r \geq 2$ ) infinite binary sequences that is  $f_{ij} \in \{0, 1\}$  ( $i = 1, \dots, r$ ;  $j = 1, 2, \dots$ ).

We process these sequences using the following algorithm [1].

1. Processing proceeds in the discrete points of time 1, 2, .... Let  $t = 1$ .
2. Let  $B_t = (f_{1t}, f_{2t}, \dots, f_{rt})$ .
3. In the moment of time  $t$  let us distinguish three cases:

- a)  $f_{12} \neq f_{11}$ ;
- b) there exists an index  $k$  ( $1 \leq k < r$ ) for which  $f_{12} = f_{11} = f_{21} = \dots = f_{k1}$  and  $f_{k+1,1} \neq f_{11}$ ;
- c)  $f_{12} = f_{11} = f_{21} = \dots = f_{r1}$ .

In the case a) the elements  $f_{11}$  and  $f_{12}$ , in the case b) the elements  $f_{11}$  and  $f_{k+1,1}$ , and in the case c) only the element  $f_{11}$  will be processed.

4. We omit the processed elements, reduce the second index of the remaining elements in the  $i$ -th sequence by the number of the omitted ones in the  $t$ -th point of time elements of the  $i$ -th ( $i = 1, \dots, r$ ) sequence.

5. We add 1 to  $t$  and continue the processing from the Step 2.

Of course, the element  $f_{i1}$  of  $B_t$  ( $t = 1, 2, \dots$ ) is the starting (first non-processed) element of  $F_i$  in the  $t$ -th point of time; we call  $B_1, B_2, \dots$  the state sequence of processing.

Let now

$$(2.2) \quad \begin{aligned} \theta_1 &= \xi_{11}, \xi_{12}, \dots \\ &\vdots \\ \theta_r &= \xi_{r1}, \xi_{r2}, \dots \end{aligned}$$

be  $r$  sequences of independent random variables with common distributions

$$(2.3) \quad P(\xi_{ij} = 0) = P(\xi_{ij} = 1) = \frac{1}{2} \quad (i = 1, \dots, r; j = 1, 2, \dots).$$

Let us suppose that the realisations of the sequences are processed using the algorithm given above.

If the vectors  $\sigma_t$  ( $t = 1, 2, \dots$ ) are determined by the distribution of the possible values of  $B_t$ 's, then  $\sigma_t$ 's are random variables.

Our algorithm and scheme is a special case of those studied in [2]; in that work ergodicity was proved for the studied systems.

In [3] the variables  $\xi_{ij}$  are independent, uniformly distributed on the set  $\{1, \dots, N\}$ , and the question was considered: how many rows are needed to process  $N$  (different) elements? Limit distribution was given for this random variable for  $t = 1, N \rightarrow \infty$ . In [5] limit distribution is given for it if  $t > 1$  fixed and  $N \rightarrow \infty$ .

For any such system the problem of the speed, especially the asymptotic speed (when  $t \rightarrow \infty$ ) of the algorithm is of basic importance, as Katai pointed out [4]. The asymptotic speed depends on the ergodic behaviour of  $\sigma_t$ . As we mentioned already,  $\sigma_t$  is an ergodic Markov-chain, and to derive its ergodic distribution is our next aim in the case  $N = 2, p = \frac{1}{2}$ .

First of all let us compute the transition probabilities for this sequence in the case  $r = 3$ .

Then there are  $2^3 = 8$  possible vectors:  $(0, 0, 0), (0, 0, 1), \dots, (1, 1, 1)$ . For the sake of simplicity we denote the vector  $(x_1, x_2, x_3)$  by

$$(2.4) \quad V = 4x_1 + 2x_2 + x_3, \quad V = 0, 1, \dots, 7 \quad (x_i = 0, 1).$$

The transition probabilities (only the positive ones) are presented in the following table.

Table 1.

| Transition probabilities for $r = 3$ |     |     |     |     |     |     |     |     |
|--------------------------------------|-----|-----|-----|-----|-----|-----|-----|-----|
|                                      | 0   | 1   | 2   | 3   | 4   | 5   | 6   | 7   |
| 0                                    | 3/4 | —   | —   | —   | 1/4 | —   | —   | —   |
| 1                                    | 1/4 | 2/4 | —   | —   | —   | 1/4 | —   | —   |
| 2                                    | 1/4 | —   | 2/4 | —   | —   | —   | 1/4 | —   |
| 3                                    | —   | 1/4 | —   | 2/4 | —   | —   | —   | 1/4 |
| 4                                    | 1/4 | —   | —   | —   | 2/4 | —   | 1/4 | —   |
| 5                                    | —   | 1/4 | —   | —   | 2/4 | —   | 1/4 | —   |
| 6                                    | —   | —   | 1/4 | —   | —   | —   | 2/4 | 1/4 |
| 7                                    | —   | —   | 1/4 | —   | —   | —   | —   | 3/4 |

We get these probabilities as follows. For the vectors 0 and 7 we have  $f_{11} = f_{21} = \dots = f_{r1}$ , therefore only the cases *a)* and *c)* are possible in Step 3. of the algorithm. With probability  $P(f_{11} \neq f_{12} \text{ and } f_{12} = f_{13}) = \frac{1}{4}$  we get the transitions  $0 \rightarrow 4$  or  $7 \rightarrow 3$  and with

$$(2.5) \quad P(f_{11} = f_{12}) + P(f_{11} \neq f_{12} \text{ and } f_{13} = f_{11}) = \frac{1}{2} + \frac{1}{4} = \frac{3}{4}$$

we have the transitions  $0 \rightarrow 0$  or  $7 \rightarrow 7$ , that is

$$(2.6) \quad P(0 \rightarrow 4) = P(7 \rightarrow 3) = \frac{1}{4} \quad \text{and} \quad P(0 \rightarrow 0) = P(7 \rightarrow 7) = \frac{3}{4}.$$

For the remaining vectors only the cases *a)* and *b)* are possible.

Let

$$(2.7) \quad \bar{f}_{ij} = \begin{cases} 1, & \text{if } f_{ij} = 0, \\ 0, & \text{if } f_{ij} = 1, \end{cases}$$

and

$$(2.8) \quad \overline{f_1, f_2, \dots, f_r} = \bar{f}_1, \bar{f}_2, \dots, \bar{f}_r.$$

Then for the vector  $s$  ( $1 \leq s \leq 6$ ) let  $f_{11} = f_{21} = \dots = f_{k1}$  and  $f_{k+1,1} \neq f_{11}$ . Now

$$(2.9) \quad P(s \rightarrow s) = P(f_{11} = f_{12} \text{ and } f_{k+1,1} = f_{k+1,2}) + \\ + P(f_{11} \neq f_{12} \text{ and } f_{13} = f_{11}) = \frac{2}{4},$$

$$(2.10) \quad \begin{aligned} P((f_{11}, \dots, f_{r1}) \rightarrow (f_{11}, f_{21}, \dots, f_{r1})) &= \\ &= P(f_{11} \neq f_{12} \text{ and } f_{13} = f_{12}) = \frac{1}{4}, \end{aligned}$$

and

$$\begin{aligned} P((f_{11}, \dots, f_{k,1}, f_{k+1,1}, \dots, f_{r1}) \rightarrow (f_{11}, \dots, f_{k,1}, f_{k+1,1}, \dots, f_{r1})) &= \\ &= P(f_{11} = f_{12} \text{ and } f_{k+1,1} \neq f_{k+1,2}) = \frac{1}{4}. \end{aligned}$$

It is easy to see that the sequence  $\sigma_1, \sigma_2, \dots$  is a homogeneous ergodic Markov-chain.

If

$$\sigma_t = (f_{11}, \dots, f_{r1}),$$

and

$$(2.11) \quad \begin{aligned} f_{11} = f_{12} = \dots = f_{1,r+1} &= 0, \\ f_{22} = f_{32} = \dots = f_{r2} &= 0, \end{aligned}$$

then  $\sigma_{t+r} = (0, 0, \dots, 0)$ , that is from any vector  $\sigma_t$  we can reach the vector  $(0, 0, \dots, 0)$  in  $r$  transitions with a probability which is not less than  $\left(\frac{1}{2}\right)^{2r}$ , and this fact is enough to guarantee ergodicity due to the Markov-theorem.

In the following paragraph we shall determine the ergodic distribution, that is the limit probabilities

$$\lim_{t \rightarrow \infty} P(\sigma_t = j) = p_j \quad (j = 0, 1, \dots, 2^r - 1).$$

### §. 3. The ergodic probabilities

Now we are going to determine the limit distribution of the vectors of the first non-processed elements.

If we process  $r$  sequences, then the vectors have  $2^r$  possible values. The ergodic probabilities have to satisfy the following equations:

$$(3.1) \quad p_j = \sum_{i=1}^n p_i p_{ij} \quad (j = 0, 1, \dots, 2^r - 1)$$

and

$$(3.2) \quad \sum_{j=0}^{2^r-1} p_j = 1.$$

Before the general case let us consider the case  $r = 3$ , whose transition probabilities are given in Table 1.

In this case

$$(3.3) \quad p_0 = \frac{3}{4}p_0 + \frac{1}{4}p_1 + \frac{1}{4}p_2 + \quad + \frac{1}{4}p_4, \quad ,$$

$$(3.4) \quad p_1 = \quad \frac{2}{4}p_1 + \quad + \frac{1}{4}p_3 + \quad + \frac{1}{4}p_5, \quad ,$$

$$(3.5) \quad p_2 = \quad \frac{2}{4}p_2 + \quad + \frac{1}{4}p_6, \quad ,$$

$$(3.6) \quad p_3 = \quad + \frac{2}{4}p_3 + \quad + \frac{1}{4}p_7, \quad ,$$

$$(3.7) \quad p_4 = \frac{1}{4}p_0 + \quad + \frac{2}{4}p_4, \quad ,$$

$$(3.8) \quad p_5 = \quad \frac{1}{4}p_1 + \quad + \frac{2}{4}p_5, \quad ,$$

$$(3.9) \quad p_6 = \quad \frac{1}{4}p_2 + \quad + \frac{1}{4}p_4 + \quad + \frac{2}{4}p_6, \quad ,$$

$$(3.10) \quad p_7 = \quad + \frac{1}{4}p_3 + \quad + \frac{1}{4}p_5 + \frac{1}{4}p_6 + \frac{3}{4}p_7, \quad ,$$

and

$$(3.11) \quad p_0 + p_1 + p_2 + p_3 + p_4 + p_5 + p_6 + p_7 = 1.$$

In consequence of the symmetry we have

$$(3.12) \quad p_7 = p_0, \quad p_6 = p_1, \quad p_5 = p_2 \quad \text{and} \quad p_4 = p_3,$$

and so instead of the system (3.3)–(3.11) we get

$$(3.13) \quad 4p_1 = 2p_1 + p_2 + p_3,$$

$$(3.14) \quad 4p_2 = p_1 + 2p_2, \quad ,$$

$$(3.15) \quad 4p_3 = p_0 + 2p_3,$$

and

$$(3.16) \quad 2p_0 + 2p_1 + 2p_2 + 2p_3 = 1.$$

(We remark, that (3.3) was omitted because of the redundancy.)

From (3.14) and (3.15) we get

$$(3.17) \quad 2p_2 = p_1, \quad 2p_3 = p_0.$$

Substituting (3.17) into (3.16) we get

$$(3.18) \quad 3p_0 + 3p_1 = 1.$$

Substituting (3.17) into (3.13)

$$(3.19) \quad 3p_1 = p_0.$$

Substituting (3.19) into (3.18)

$$(3.20) \quad 4p_0 = 1,$$

and then

$$(3.21) \quad \begin{aligned} p_0 = p_7 = \frac{1}{4}, \quad p_1 = p_6 = \frac{1}{4 \cdot 3}, \\ p_2 = p_5 = \frac{1}{4 \cdot 3 \cdot 2}, \quad p_3 = p_4 = \frac{1}{4 \cdot 2}. \end{aligned}$$

Generalizing the previous computations we find a connection among the  $p$ 's.

Let  $A_j$  denote any sequence of  $j$  ( $1 \leq j \leq r$ ) binary digits, that is let

$$(3.22) \quad A_j = i_1, i_2, \dots, i_j \quad (i_1, \dots, i_j \in \{0, 1\}).$$

We shall use also the notation

$$(3.23) \quad \bar{A}_j = \bar{i}_1, \bar{i}_2, \dots, \bar{i}_j.$$

Let  $\alpha_r(A_r)$  denote the ergodic probability of the vector  $A_r$  in the case of  $r$  processable sequences.

At first we prove the following assertion.

**Lemma 1.** *If  $r \geq 2$ ,  $1 \leq k \leq r$ , then*

$$(3.24) \quad k \alpha_r(00 \dots 001 A_{r-k}) = \alpha_r(00 \dots 000 A_{r-k}). \quad \square$$

According to this assertion the ergodic probability of a binary vector consisting of  $(k-1)$  zeros, one 1 and any sequence  $A_{r-k}$  of length  $(r-k)$  is  $k$  times smaller than the ergodic one of the binary vector in which  $k$  ones are continued by the sequence  $A_{r-k}$ .

**Proof of Lemma 1.**

For  $k = 1$  we assert

$$(3.25) \quad \alpha_r(1 A_{r-1}) = \alpha_r(0 A_{r-1}),$$

which holds in consequence of the symmetry  $\overline{1 A_{r-1}} = \overline{0 A_{r-1}}$ .

Let us suppose, (3.24) holds for  $k = 1, \dots, s$ , where  $s < r$ . Then we show that (3.25) holds for  $k = s+1$ , too.

Let us consider the binary vectors

$$(3.26) \quad 00 \dots 001 A_{r-s-1}.$$

We get this vector with a probability  $\frac{2}{4}$  from itself, and with a probability  $\frac{1}{4}$  from the vectors, whose first  $s$  elements contain punctually one 1, and the last  $r-s$  elements are  $1 A_{r-s-1}$ .

Let  $B_j$  ( $j = 1, \dots, s$ ) be the binary vector of length  $s$ , in which the  $j$ -th element is 1, the remaining elements are 0's.

Then

$$(3.27) \quad \begin{aligned} & \alpha_r(00 \dots 001 A_{r-s-1}) = \\ & = \frac{2}{4} \alpha_r(00 \dots 001 A_{r-s-1}) + \sum_{j=1}^s \frac{1}{4} \alpha_r(B_j 1 A_{r-s-1}), \end{aligned}$$

and from here

$$(3.28) \quad 2\alpha_r(00 \dots 001 A_{r-s-1}) = \sum_{j=1}^s \alpha_r(B_j 1 A_{r-s-1}).$$

Using twice the induction hypothesis we get for  $j = 1, \dots, s-1$

$$(3.29) \quad \begin{aligned} & \alpha_r(B_j 1 A_{r-s-1}) = \alpha_r(00 \dots 010 \dots 01 A_{r-s-1}) = \\ & = \frac{1}{j} \alpha_r(00 \dots 001 \dots 10 A_{r-s-1}) = \\ & = \frac{1}{(j+1)j} \alpha_r(00 \dots 000 \dots 01 \overline{A_{r-s-1}}). \end{aligned}$$

For  $j = s$  the induction hypothesis gives

$$(3.30) \quad \begin{aligned} & \alpha_r(00 \dots 000 \dots 11 A_{r-s-1}) = \\ & = \frac{1}{s} \alpha_r(00 \dots 000 \dots 00 \overline{A_{r-s-1}}). \end{aligned}$$

Let us substitute (3.29) and (3.30) into (3.28):

$$(3.31) \quad \begin{aligned} & \alpha_r(00 \dots 001 A_{r-s-1}) \left[ 2 - \sum_{j=1}^{s-1} \frac{1}{j(j+1)} \right] = \\ & = \frac{1}{s} \alpha_r(00 \dots 00 \overline{A_{r-s-1}}). \end{aligned}$$

Taking into account

$$(3.32) \quad \frac{1}{j(j+1)} = \frac{1}{j} - \frac{1}{j+1},$$

we get

$$(3.33) \quad \left(1 + \frac{1}{s}\right) \alpha_r(00 \dots 001 A_{r-s-1}) = \frac{1}{s} \alpha_r(00 \dots 00 \overline{A_{r-s-1}}),$$

that is

$$(3.34) \quad (s+1) \alpha_r(00 \dots 001 A_{r-s-1}) = \alpha_r(00 \dots 00 \overline{A_{r-s-1}}). \quad \square$$

Now we compute the ergodic probability of the vectors consisting of identical elements.

**Lemma 2.** *If  $r \geq 2$  then*

$$(3.35) \quad \alpha_r(00 \dots 00) = \alpha_r(11 \dots 11) = \frac{1}{r+1}. \quad \square$$

**Proof.** We begin with the equality

$$(3.36) \quad \sum_{j=0}^{2^r-1} p_j = 1.$$

In consequence of the symmetry we have for  $j = 2^{r-1}, \dots, 2^r - 1$

$$(3.37) \quad p_j = p_{2^{r-1}-j},$$

and therefore

$$(3.38) \quad \sum_{j=0}^{2^r-1} p_j = \sum_{j=0}^{2^{r-1}-1} p_j + \sum_{j=2^{r-1}}^{2^r-1} p_j = 2 \sum_{j=0}^{2^{r-1}-1} p_j.$$

We have got that for  $k = 1$  holds

$$(3.39) \quad (k+1) \sum_{j=0}^{2^{r-k}-1} p_j = 1.$$

Let us suppose that (3.39) holds for  $k = s$ , where  $s < r$ . We show that than (3.39) holds for  $k = s+1$  too.

By the induction hypothesis we have

$$(3.40) \quad (s+1) \sum_{j=0}^{2^{r-s}-1} p_j = 1.$$

We divide the numbers  $0, 1, \dots, 2^{r-s}-1$  into two groups:  $0, 1, \dots, 2^{r-s-1}-1$  and  $2^{r-s-1}, \dots, 2^{r-s}-1$ . The elements of the groups have the form  $00 \dots 00 A_{r-s-1}$  and  $00 \dots 01 \overline{A_{r-s-1}}$  resp. and we have a

one to one correspondence among them (the sum of the corresponding elements equal to  $2^{r-s}-1$ .) Therefore by Lemma 1 we get

$$(3.41) \quad (s+1) \sum_{j=0}^{2^{r-s}-1} p_j = (s+1) \sum_{j=0}^{2^{r-s-1}-1} p_j + (s+1) \sum_{j=2^{r-s-1}}^{2^{r-s}-1} p_j =$$

$$(3.42) \quad = (s+1) \sum_{j=0}^{2^{r-s-1}-1} p_j + (s+1) \sum_{j=0}^{2^{r-s-1}-1} \frac{p_j}{s+1} = (s+2) \sum_{j=0}^{2^{r-s-1}-1} p_j.$$

Hence (3.42) holds for every  $k = 1, \dots, r$ , among them for  $k = r$ , therefore

$$(3.43) \quad (r+1) p_0 = 1.$$

From (3.43) and the symmetry we get (3.35).  $\square$

Now we can determine the ergodic probabilities.

**Theorem 1.** *Let us process the binary sequences  $\xi_{i1}, \xi_{i2}, \dots$ , ( $i = 1, \dots, r$ ) of independent random variables with common distribution*

$$(3.44) \quad P(\xi_{ij} = 0) = P(\xi_{ij} = 1) = \frac{1}{2} \quad (i = 1, \dots, r; j = 1, 2, \dots)$$

*using the algorithm defined above. Then the ergodic probability of the vector  $(i_0, \dots, i_{r-1})$  equals to*

$$(3.45) \quad \alpha_r(i_0, \dots, i_{r-1}) = \frac{1}{(r+1) \prod_{\substack{2 \leq k \leq r \\ i_{k-1} \neq i_k}} k} \cdot \square$$

According to this theorem we get  $\alpha_r(i_0, \dots, i_{r-1})$  dividing the basic value  $\alpha_r(00 \dots 00) = \frac{1}{r+1}$  by the product of the indices of  $i$ 's differing from the previous one.

**Proof of theorem 1.** Let

$$(3.46) \quad j = \sum_{k=0}^{r-1} i_k \cdot 2^k$$

be the binary form of  $j$  for  $j = 0, 1, \dots, 2^r - 1$ . Due to the symmetry we have to deal only with the values  $j = 0, 1, \dots, 2^{r-1} - 1$ .

If  $r = 1$ , then a simple direct calculation gives  $\alpha_1(0) = \frac{1}{2}$ . Let now  $r \geq 2$  be.

As for  $j = 0$  the product in (3.45) is empty (i.e. there is no a change in the corresponding binary vector), Lemma 2. gives the same value, as (3.45).

Let now be  $s$  ( $1 \leq s \leq r-1$ ) changes in  $j$ , that is let differ punctually the elements on the  $h_1$ -th,  $\dots$ ,  $h_s$ -th places from the elements preceeding them.

Using Lemma 1 for  $k = h_1, h_2, \dots, h_s$ , we get

$$(3.47) \quad \alpha_r(j) \prod_{i=1}^s h_i = \alpha_r(0).$$

That was to be proved.  $\square$

#### §. 4. On the processing speed

Processing binary sequences as before let the random variable  $\eta_{ij}$  denote the number of the processed elements of the  $i$ -th sequence in the  $j$ -th point of time.

According to the processing algorithm we have

$$(4.1) \quad 1 \leq \eta_{1j} \leq 2 \quad (j = 1, 2, \dots),$$

$$(4.2) \quad 0 \leq \eta_{ij} \leq 1 \quad (i = 2, \dots, r; j = 1, 2, \dots),$$

$$(4.3) \quad 1 \leq \sum_{i=1}^r \eta_{ij} \leq 2 \quad (j = 1, 2, \dots).$$

The processing speed  $S^{(r)}$  is defined by

$$(4.4) \quad S^{(r)} = \lim_{t \rightarrow \infty} \frac{\sum_{j=1}^t M \left( \sum_{i=1}^r \eta_{ij} \right)}{t}.$$

There is a close connection between the distributions of  $\sigma_j$  and  $\sum_{i=1}^r \eta_{ij}$ :

$$(4.5) \quad \begin{aligned} P \left( \sum_{i=1}^r \eta_{ij} = 1 \right) &= \frac{1}{2} P(\sigma_j = (0, 0, \dots, 0, 0)) + \\ &+ \frac{1}{2} P(\sigma_j = (1, 1, \dots, 1, 1)). \end{aligned}$$

As  $\sigma_1, \sigma_2, \dots$  is ergodic, the right side of (4.5) has a limit, whose value is  $\frac{1}{r+1}$  according to Lemma 2. Therefore

$$(4.6) \quad \lim_{j \rightarrow \infty} M \left( \sum_{i=1}^r \eta_{ij} \right) = \frac{1}{r+1} \cdot 1 + \frac{r}{r+1} \cdot 2 = 2 - \frac{1}{r+1},$$

and from where

$$(4.7) \quad S^{(r)} = 2 - \frac{1}{r+1},$$

as the convergence of the sequence (4.6) implies the convergence (to the same limit) of the middle values of its elements in (4.4).

If we define the processing speed  $S_i$  for the  $i$ -th sequence by

$$(4.8) \quad S_1 = S^{(1)} \quad \text{and} \quad S_i = S^{(i)} - S^{(i-1)} \quad (i = 2, \dots, r),$$

then from (4.7) we get

$$(4.9) \quad S_1 = \frac{3}{2} \quad \text{and} \quad S_i = \frac{1}{i(i+1)} \quad (i = 2, \dots, r).$$

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### REFERENCES

- [1] A. Iványi, I. Kátaí: Processing of random sequences with priority, *Acta Cybernetica*, 4 (1975), 85–101.
- [2] A. Iványi, I. Kátaí: Processing of independent Markov-chains, *Annales Univ. Budapestinensis, Sectio Computatorica*, 3 (1982), 33–46.
- [3] A. Iványi, I. Kátaí: Parallel processing of random sequences with priority, *Lecture Notes Stat.*, 8 (1981), 122–134.
- [4] А. Ивани, И. Катаи: Моделирование бесприоритетного режима ЭВМ с блочной памятью при полной информации, *Автоматика и телемеханика*, 39 (1984).
- [5] Pergel, J.: The lowest priority in processing of random sequences (to appear).

