A PROOF OF THE GENERALIZED HADAMARD INEQUALITY VIA INFORMATION THEORY

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The generalized Hadamard inequality is considered in this paper. The new proof presented here is motivated by the idea of mutual information of random variables. For a strict algebraic proof see e.g. [3].

Let T be and $m \times n$ matrix of real numbers and denote its column vectors with t_1, t_2, \ldots, t_n . Let the set of column indices $J = \{1, 2, \ldots, n\}$ partitioned into sets J_1, J_2, \ldots, J_ρ , that is $J_i \cap J_j = \emptyset$ for $i \neq j$, $\bigcup_{k=1}^{\rho} J_k = J$. Let $T_k = \{t_j : j \in J_k\}$. Assume that the vectors $t_j \in T_k$ are linearly independent for $k = 1, 2, \ldots, \rho$. Now the generalized Hadamard inequality can be stated as follows.

Theorem.

$$\det(T^*T) \leq \det(T_1^*T_1)\det(T_2^*T_2) \dots \det(T_\rho^*T_\rho)$$

and equality holds iff $T_i^*T_j = 0$ for all $i \neq j$. (A* stands for the transpose of the matrix A).

Note that if $\rho = n$ and $T_1 = \{t_1\}, T_2 = \{t_2\}, \ldots, T_n = \{t_n\}$ then the assumption says that $t_j \neq 0$ for all j and the theorem states the well-known version of the Hadamard inequality, that is

$$\det(T^*T) \le t_1^2 t_2^2 \dots t_n^2$$

and equality holds iff $t_i^*t_j = 0$ for all $i \neq j$.

Our proof goes through a couple of remarks.

Remark 1. Consider a real-valued function $f(x), x = (x_1, x_2, \ldots, x_n)$. Let the set of indices $J = \{1, 2, \ldots, n\}$ be partitioned into sets J_1, \ldots, J_ρ . After a proper rearranging of the components x_1, x_2, \ldots, x_n we can split the vector x into subvectors x_1, x_2, \ldots, x_ρ such that x_k contains the components $x_j, j \in J_k$. With this notation $f(x) = f(x_1, x_2, \ldots, x_\rho)$. Assume that

(i)
$$f(x_1, x_2, ..., x_\rho) \ge 0$$
 $(x_1, x_2, ..., x_\rho) \in \mathbb{R}^n$,

(ii)
$$\int_{-\infty}^{\infty} \dots \int_{-\infty}^{\infty} f(x_1, x_2, \dots, x_{\rho}) dx_1 dx_2 \dots dx_{\rho} = 1.$$

Assumptions (i), (ii) obviously hold for a density function of an *n*-dimensional random variable. Introducing the functions

$$f_k(x_k) = \int_{-\infty}^{\infty} \dots \int_{-\infty}^{\infty} f(x_1, x_2, \dots, x_{\rho}) dx_1 \dots dx_{k-1} dx_{k+1} \dots dx_{\rho}, \quad k = 1, 2, \dots, \rho$$

we have

$$\int_{-\infty}^{\infty} \dots \int_{-\infty}^{\infty} f_k(x_k) dx_k = 1 \quad k = 1, 2, \dots, \rho$$

and

$$\int_{-\infty}^{\infty} \ldots \int_{-\infty}^{\infty} f(x_1, x_2, \ldots, x_{\rho}) \log \frac{f(x_1, x_2, \ldots, x_{\rho})}{f_1(x_1) f_2(x_2) \ldots f_{\rho}(x_{\rho})} dx_1 dx_2 \ldots dx_{\rho} \geq 0.$$

Here equality holds iff $f(x_1, x_2, ..., x_\rho) = f_1(x_1) f_2(x_2) ... f_\rho(x_\rho)$. The latter inequality plays an important role in our discussion and is well-known in information theory (Shannon lemma). It can be proved from the fact that

$$\log t \ge 1 - \frac{1}{t} \quad \text{for } t \ge 0$$

and equality holds iff t = 1.

Remark 2. Consider a matrix T for which all assumptions incurred in the theorem are valid. Then the $n \times n$ matrix T^*T is positive definite, symmetric and can be partitioned as follows.

$$T^*T = \begin{bmatrix} T_1^*T_1 & T_1^*T_2 & \dots & T_1^*T_{\rho} \\ T_2^*T_1 & T_2^*T_2 & \dots & T_2^*T_{\rho} \\ \vdots & \vdots & & \vdots \\ T_{\rho}^*T_1 & T_{\rho}^*T_2 & \dots & T_{\rho}^*T_{\rho} \end{bmatrix}$$

With this T^*T let us consider the function

$$g(x) = \frac{1}{\pi^{n/2}} \frac{1}{(\det(T^*T))^{1/2}} \exp\left(-x^*(T^*T)^{-1}x\right).$$

g(x) is the density function of an *n*-dimensional random variable with Gauss distribution. Obviously

$$g(x) \ge 0$$
 $x \in \mathbb{R}^n$

$$\int_{-\infty}^{\infty} \dots \int_{-\infty}^{\infty} g(x)dx = 1.$$

Let us partition the components of x into subvectors $x_1, x_2, \ldots, x_{\rho}$ according to the partition $J_1, J_2, \ldots, J_{\rho}$ of the column indices $\{1, 2, \ldots, n\}$. One can show by evaluating the integral (see e.g.[1]) that

$$g_k(x_k) = \int\limits_{-\infty}^{\infty} \dots \int\limits_{-\infty}^{\infty} g(x_1, x_2, \dots, x_{
ho}) dx_1 \dots dx_{k-1} dx_{k+1} \dots dx_{
ho} =$$

$$= \frac{1}{\pi^{|J_k|/2}} \frac{1}{(\det(T_k^* T_k))^{1/2}} \exp\left(-x_k^* (T_k^* T_k)^{-1} x_k\right)$$

where $|J_k|$ denotes the cardinality of the set J_k .

Remark 3. Evaluating the integrals one can easily check that

$$\int_{-\infty}^{\infty} \dots \int_{-\infty}^{\infty} g(x_1, x_2, \dots, x_{
ho}) \log \frac{g(x_1, x_2, \dots, x_{
ho})}{g_1(x_1)g_2(x_2) \dots g_{
ho}(x_{
ho})} dx_1 dx_2 \dots dx_{
ho} =$$

$$= \int_{-\infty}^{\infty} \dots \int_{-\infty}^{\infty} g(x_1, x_2, \dots, x_{
ho}) \log g(x_1, x_2, \dots, x_{
ho}) dx_1 dx_2 \dots dx_{
ho} -$$

$$- \sum_{k=1}^{\rho} \left(\int_{-\infty}^{\infty} \dots \int_{-\infty}^{\infty} g_k(x_k) \log g_k(x_k) dx_k \right) =$$

$$= \frac{1}{2} \log \frac{\det(T_1^*T_1) \det(T_2^*T_2) \dots \det(T_{
ho}^*T_{
ho})}{\det(T^*T)}.$$

Now applying the inequality of Remark 1 to g(x) states that the integral above is non-negative, and equals zero iff

	$T_{1}^{*}T_{1}$	$T_{1}^{*}T_{2}$		$T_1^*T_{ ho}$		$T_1^*T_1$	0	 0
	$T_2^*T_1$	$T_2^*T_2$		$T_2^*T_ ho$		0	$T_2^*T_2$	 0
	:	:		:	=	:	:	:
-	$T_{ ho}^*T_1$	$T_{ ho}^*T_2$	• • • •	$T_{\rho}^{*}T_{\rho}$		0	0	 $T_{\rho}^{*}T_{\rho}$

Hence the theorem is proved.

Corollary. Let us consider a positive definite symmetric matrix C of size $n \times n$. Take a partition of indices $J = \{1, 2, ..., n\}$ into the sets $J_1, J_2, ..., J_{\rho}$ and denote the submatrix of elements c_{ij} $i \in J_k, j \in J_l$ with C_{kl} . Then

$$\det(C) \leq \det(C_{11})\det(C_{22}) \dots \det(C_{\rho\rho})$$

and equality holds iff $C_{kl} = 0$ for all $k \neq l$. To prove the corollary it should be noted that according to Cholesky's factorization theorem for every positive definite, symmetric matrix $C = T^*T$ with a suitable non-singular matrix T. Then the theorem directly applies.

References

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